

# EXACT BEHAVIOR OF GAUSSIAN SEMINORMS

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## ABSTRACT

The exact lower tail of Gaussian seminorms are evaluated, using a refinement of the techniques presented in [5].

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**I. Introduction** Let  $\{x_i\}_{i=1}^{\infty}$  be a sequence of independent, standard Gaussian random variables. Consider the random variable  $z = \sum_{i=1}^{\infty} x_i^2/a_i^2$ , where  $\{a_i\}_{i=1}^{\infty}$  is a sequence of given (deterministic) finite numbers satisfying  $\sum_{i=1}^{\infty} \frac{1}{a_i^2} < \infty$ . We are interested in computing the asymptotics  $P(z < \epsilon)$  as  $\epsilon \rightarrow 0$ . This problem was considered in the 70's by the Soviet school [6, 7, 3]. Their approach was analytical in flavor and based on the saddle point method. Independently, unaware of this, a probabilistic study of  $P(z < \epsilon)$  was initiated by [2] and improved upon in [5], the latter using large deviations techniques. This study provides explicit and easy bounds on  $P(z < \epsilon)$  which, however, are not asymptotically tight. The purpose of this note is to push this analysis to yield a tight asymptotic evaluation of the above probability. We thus retrieve, by probabilistic methods, the analytic results of [6, 7, 3], in a shorter, more transparent way, which seems amenable to extensions in different directions.

To state precisely our goal, recall that the results in [5] imply that there exist functions

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$g(\cdot)$  and  $I(\cdot)$  such that  $z/\epsilon$  satisfies the large deviations principle with rate  $g(\epsilon)$  and rate function  $I(\cdot)$ . In particular, it is shown there that for some constant  $c > 0$ , which depends on the sequence  $\{a_i\}$  and may be computed explicitly,

$$\lim_{\epsilon \rightarrow 0} g(\epsilon) \log P(z < \epsilon) = -c. \quad (1)$$

We show below that one may find an explicit function  $f(\epsilon)$  such that

$$\lim_{\epsilon \rightarrow 0} f(\epsilon) P(z < \epsilon) = 1 \quad (2)$$

(see (10)). For related results and comparison theorems for small balls probabilities see [4].

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**II. Computation of Asymptotic Probabilities** As in the Introduction, let  $\{x_i\}$  be i.i.d. standard Gaussian variables, defined on a suitable probability space  $(\Omega, \mathcal{F}, P)$ . For  $\theta \geq 0$  let  $\Lambda(\theta) \triangleq \log E(e^{-\theta x_1^2}) = -\frac{1}{2} \log(1 + 2\theta)$  and define

$$\mu(\theta) \triangleq -\sum_{i=1}^{\infty} \frac{1}{a_i^2} \Lambda'\left(\frac{\theta}{a_i^2}\right) = \sum_{i=1}^{\infty} \frac{1}{a_i^2 + 2\theta} \quad (3)$$

$$\psi(\theta) \triangleq \sqrt{\sum_{i=1}^{\infty} \frac{\theta^2}{a_i^4} \Lambda''\left(\frac{\theta}{a_i^2}\right)} = \sqrt{\sum_{i=1}^{\infty} \frac{2\theta^2}{(a_i^2 + 2\theta)^2}} \quad (4)$$

$$I(\theta) \triangleq \sum_{i=1}^{\infty} \left( \frac{\theta}{a_i^2} \Lambda'\left(\frac{\theta}{a_i^2}\right) - \Lambda\left(\frac{\theta}{a_i^2}\right) \right) = \frac{1}{2} \sum_{i=1}^{\infty} \log\left(1 + \frac{2\theta}{a_i^2}\right) - \theta \mu(\theta) \quad (5)$$

Under our assumptions, there exists a unique  $\theta_\epsilon$  such that  $\mu(\theta_\epsilon) = \epsilon$ ,  $\theta_\epsilon \rightarrow_{\epsilon \rightarrow 0} \infty$ , and  $\psi(\theta_\epsilon) \rightarrow_{\epsilon \rightarrow 0} \infty$ . In the sequel, we use  $\theta, \mu, \psi, I$  for  $\theta_\epsilon, \mu(\theta_\epsilon), \psi(\theta_\epsilon), I(\theta_\epsilon)$ . Define on  $(\Omega, \mathcal{F})$  a probability measure by

$$dP_\theta = \frac{e^{-\theta z} dP}{E_P(e^{-\theta z})} \quad (6)$$

Then

$$P(z < \epsilon) = e^{-I} \int_{\Omega} e^{\theta(z-\mu)} 1_{\{0 \leq z < \mu\}} dP_\theta \quad (7)$$

Defining now

$$U_\psi = \frac{\theta(\mu - z)}{\psi},$$

and denoting by  $\eta_\psi$  the law of  $U_\psi$  under  $P_\theta$ , one obtains

$$e^I P(z < \epsilon) = \int_{\Omega} e^{-\psi U_\psi} 1_{\{0 < U_\psi \leq \theta\mu/\psi\}} dP_\theta = \int_0^{\mu\theta/\psi} e^{-\psi v} d\eta_\psi(v) = \int_0^\infty e^{-\psi v} d\eta_\psi(v) \quad (8)$$

where we used the fact that  $\eta_\psi$  is supported on  $(-\infty, \theta\mu/\psi]$ . Integrating by parts, one obtains

$$e^I P(z < \epsilon) = \int_0^\infty e^{-v} \eta_\psi([0, v/\psi]) dv. \quad (9)$$

It is straightforward to check that  $E_{\eta_\psi}(U_\psi) = 0$  and  $E_{\eta_\psi}(U_\psi^2) = 1$ . Thus, if we knew that for  $\epsilon \rightarrow 0$ ,  $\eta_\psi$  converges locally (in the scale  $1/\psi$ ) to the standard Gaussian law, one would conclude from (9) the following

**Theorem:**

$$\lim_{\epsilon \rightarrow 0} e^I \psi P(z < \epsilon) = 1/\sqrt{2\pi}. \quad (10)$$

**Remark:** The behavior of  $I$  in  $\epsilon$  may be read off its definition, and in many cases can be made explicit (see, e.g., displays (16) and (17) in [5]). A similar computation yields also the behavior of  $\psi$  in  $\epsilon$ .

**Proof:** Define the (infinite) triangular array

$$U_\psi = \sum_{k=1}^{\infty} b_k \xi_k, \quad (11)$$

where under  $\eta_\psi$ ,  $\{\xi_k\}_{k=1}^{\infty}$  is a  $(1 - N(0, 1)^2)$  i.i.d. sequence and

$$b_k = \frac{\theta}{\psi(a_k^2 + 2\theta)}. \quad (12)$$

We now claim that the laws  $\eta_\psi$  converge to the standard Gaussian distribution as  $\psi \rightarrow \infty$ . Indeed, since in (11)  $E_{\eta_\psi} \xi_k = 0 \quad \forall k \in \mathbb{N}$  and  $\sum_{k=1}^{\infty} b_k^2 E_{\eta_\psi} \xi_k^2 = 1$ , one only needs to prove that, for some suitable  $N(\epsilon)$ ,

$$\lim_{\psi \rightarrow \infty} \left( \sum_{k=1}^{N(\epsilon)} b_k^3 E_{\eta_\psi} |\xi_k|^3 + E_{\eta_\psi} \left| \sum_{k=N(\epsilon)+1}^{\infty} b_k \xi_k \right|^3 \right) = 0, \quad (13)$$

in order to apply classical CLT results for finite triangular arrays such as for example [1, Theorem 7.1.2]. (The condition (13) follows from expressing (11) as a finite sum). Note (it follows directly from (12)) that

$$\sum_{k=1}^{\infty} b_k^2 = \frac{1}{2}, \quad b_k \leq \frac{1}{2\psi}. \quad (14)$$

Therefore, denoting  $m_i = E|\xi_1|^i$ ,

$$\sum_{k=1}^{N(\epsilon)} b_k^3 E_{\eta_\psi} |\xi_k|^3 \leq m_3 \max_{k \in \mathbb{N}} b_k \sum_{k=1}^{\infty} b_k^2 \leq \frac{m_3}{4\psi} \xrightarrow{\psi \rightarrow \infty} 0,$$

whereas

$$E_{\eta_\psi} \left| \sum_{k=N(\epsilon)+1}^{\infty} b_k \xi_k \right|^3 \leq \left( E_{\eta_\psi} \left| \sum_{k=N(\epsilon)+1}^{\infty} b_k \xi_k \right|^4 \right)^{3/4} \leq \left( \sqrt{3m_4} \sum_{k=N(\epsilon)+1}^{\infty} b_k^2 \right)^{8/3}.$$

By taking any  $N(\epsilon) \rightarrow_{\epsilon \rightarrow 0} \infty$ , (13) follows.

It is actually possible to obtain a stronger (uniformly local) convergence. Let  $U_\psi$ 's characteristic function be denoted by  $\phi_\psi(t) = E_{\eta_\psi} e^{itU_\psi}$ .

**Lemma:** For each  $\alpha > 0$ ,

$$|\phi_\psi(t)| \leq \frac{\alpha^{\alpha/2}}{|t|^\alpha} \wedge 1 \quad \forall \psi > \alpha^{1/2}, \forall t \in \mathbb{R}. \quad (15)$$

**Proof of the Lemma:** Using the representation (11), a direct calculation yields  $\phi_\psi(t) = e^{i\frac{\theta}{\psi}\mu t} \prod_{k=1}^{\infty} (1 + 2ib_k t)^{-\frac{1}{2}}$  so that

$$|\phi_\psi(t)|^4 = \prod_{k=1}^{\infty} \frac{1}{1 + 4b_k^2 t^2} = \exp - \left\{ \sum_{k=1}^{\infty} \log(1 + 4b_k^2 t^2) \right\}. \quad (16)$$

We now use the inequality

$$\log(1 + \gamma t^2) \geq \gamma \log(1 + t^2) \quad \forall \gamma \in (0, 1), \forall t \in \mathbb{R}, \quad (17)$$

in (16). Given  $\alpha > 0$ , and if  $\psi > \alpha^{1/2}$ , then by (14)  $\gamma_k = 4\alpha b_k^2 < 1$  for all  $k \in \mathbb{N}$  and

$$\begin{aligned} |\phi_\psi(t)|^4 &\leq \exp - \left\{ \sum_{k=1}^{\infty} 4\alpha b_k^2 \log(1 + t^2/\alpha) \right\} = \exp - \{2\alpha \log(1 + t^2/\alpha)\} \\ &= \left( 1 + \frac{t^2}{\alpha} \right)^{-2\alpha} \end{aligned}$$

so that indeed  $|\phi_\psi(t)| \leq \alpha^{\alpha/2} |t|^{-\alpha}$  for all  $\psi > \alpha^{1/2}$  and all  $t \in \mathbb{R}$ . □

It then follows that

$$\lim_{\psi \rightarrow \infty} \int_{-\infty}^{\infty} \left| \phi_\psi(t) - e^{-\frac{t^2}{2}} \right| dt = 0. \quad (18)$$

Indeed, the pointwise convergence of the integrand in (18) to 0 is guaranteed by the CLT for  $\eta_\psi$ , and the  $L^1$  limit results by dominated convergence which is made possible by the Lemma (applied with any  $\alpha > 1$ ).

Denote next by  $\Gamma$  the standard Gaussian measure and by  $g(x)$  its density. From (9)

$$\left| \psi e^I P(z < \epsilon) - \frac{1}{\sqrt{2\pi}} \right| \leq \left| \psi \int_0^\infty e^{-v} \Gamma\left([0, \frac{v}{\psi}]\right) dv - \frac{1}{\sqrt{2\pi}} \right| + \psi \int_0^\infty e^{-v} |(\Gamma - \eta_\psi)\left([0, \frac{v}{\psi}]\right)| dv \triangleq J_1 + J_2.$$

Now, for some  $0 < \bar{v}_\psi < v/\psi$ ,  $J_1 = |\int_0^\infty (g(\bar{v}_\psi) - g(0))ve^{-v} dv| \xrightarrow{\psi \rightarrow \infty} 0$  by dominated convergence, while, denoting by  $f_\psi$  the density of the law  $\eta_\psi$ , whose existence follows from (15),

$$J_2 \leq \psi \int_0^\infty e^{-v} \left( \int_0^{v/\psi} |g(x) - f_\psi(x)| dx \right) dv \leq \|g - f_\psi\|_\infty \int_0^\infty ve^{-v} dv$$

which by (18) coupled with the inverse Fourier transform, converges to zero as  $\psi \rightarrow \infty$ .  $\square$

**Remarks:**

1) We mention the particular case  $a_i = i^{\beta/2}$ ,  $\beta > 1$ . A computation similar to the one in [5](displays (19), (20)) reveals that

$$I = \frac{(\beta - 1)\nu_\beta^{\frac{\beta}{\beta-1}}}{2\epsilon^{\frac{1}{\beta-1}}}(1 + o(1)), \quad \psi^2 = \frac{\nu_\beta^{\frac{1}{\beta-1}}\kappa_\beta}{2\epsilon^{\frac{1}{\beta-1}}}(1 + o(1)),$$

where  $\nu_\beta = \int_0^\infty dy/(1 + y^\beta) = \frac{\pi/\beta}{\sin(\pi/\beta)}$  and  $\kappa_\beta = \int_0^\infty dy/(1 + y^\beta)^2 = \frac{(\beta-1)/\beta^2}{\sin((\beta-1)\pi/\beta)}$ .

2) Our interest in the lower tails of the random variable  $z$  is because it represents, by the Karhunen-Loève expansion, the  $L^2$  norm of Gaussian processes. This is the reason for our considering  $x_i$  which are Gaussian random variables. The method, and perhaps the result as well, seem to depend quite heavily on the Gaussian assumption. For example, taking  $a_i^2 = 2^i$  and  $x_i$  Bernoulli random variables,  $z$  turns out to be uniform on  $[0, 1]$ , and the functions  $\mu, \psi$  and  $I$  may be computed explicitly; it turns out that  $\psi \rightarrow_{\epsilon \rightarrow 0} 1$  and the limit in (10) is  $e^{-1}$  and not  $(2\pi)^{-1/2}$ , suggesting that the central limit theorem fails in this case.

Nevertheless, under the following conditions on  $\Lambda(\theta)$  and on the sequence  $\{a_k^2\}$ , the derivation may still be carried out, defining  $\mu(\theta)$ ,  $\psi(\theta)$  and  $I(\theta)$  as in the first equality in each of (3)–(5):

**(A1)**  $\lim_{\theta \rightarrow \infty} \Lambda'(\theta) = 0$

**(A2)**  $\lim_{\theta \rightarrow \infty} \psi(\theta) = \infty$

**(A3)** (a)  $\Lambda''(y) \leq c/y^2 \ \forall y > 0$ .

(b)  $|\Lambda^{(4)}(y)| \leq c\Lambda''(y)^2 \ \forall y > 0$ .

**(A4)** for  $m$  large enough, and analytically extending  $\Lambda$  to  $\text{Re}\theta > 0$ ,

$$\Lambda(x) - \text{Re}\Lambda(x + it \frac{x}{m}) \geq (\frac{x}{m})^2 \Lambda''(x) G(t) \ \forall x > 0, t \in \mathbb{R}, \quad (19)$$

where  $G(t)$  satisfies  $\int_{-\infty}^{\infty} e^{-G(t)} dt < \infty$ .

Note that **(A1)**, **(A3)** and **(A4)** only involve  $x_1$ 's law while **(A2)** depends also on the sequence  $\{a_k^2\}$ . A sufficient condition for **(A2)** to hold for any nondegenerate  $x_1$  is that  $a_i^2$  grows subexponentially in the sense that for some  $0 < a < b < \infty$  the cardinality of  $\{i : a\theta \leq a_i^2 \leq b\theta\}$  diverges with  $\theta$ , since

$$\begin{aligned} \psi^2(\theta) &\geq \sum_{\{i: \frac{\theta}{a_i^2} \in [1/b, 1/a]\}} \frac{\theta^2}{a_i^4} \Lambda''(\frac{\theta}{a_i^2}) \\ &\geq m |\{i : \frac{1}{b} \leq \frac{\theta}{a_i^2} \leq \frac{1}{a}\}| = m |\{i : a\theta \leq a_i^2 \leq b\theta\}|, \end{aligned}$$

where  $m$  denotes the positive minimum of the function  $x^2 \Lambda''(x)$  on the interval  $[1/b, 1/a]$ .

The conditions are satisfied, for example, when  $x_1$  has a density of the form  $C_\alpha |x|^\alpha e^{-x^2/2}$ , for some  $\alpha > -1$ , since the corresponding  $\Lambda_\alpha(\theta)$  satisfy  $\Lambda_\alpha = (1 + \alpha)\Lambda_0$ . On the other hand, if  $x_1^2$  is a lattice random variable, the left hand side of (19) is periodic in  $t$  and thus **(A4)** cannot possibly hold. This is the situation, in particular, in the uniform example discussed above.

To see that these assumptions enable the above proof to be emulated, first note that **(A1)** implies, by the monotonicity of  $\Lambda'(\cdot)$ , that  $\mu(\theta) \rightarrow_{\theta \rightarrow \infty} 0$  and hence the existence of  $\theta_\epsilon$  is ensured. Next,  $U_\psi$  is represented as  $U_\psi = \sum_{k=1}^{\infty} \bar{\xi}_k$  with  $E_{\eta_\psi} \bar{\xi}_k = 0$  and  $\sum_{k=1}^{\infty} E_{\eta_\psi} \bar{\xi}_k^2 = 1$ , the third moment condition in (13) is replaced by Hölder's inequality to a fourth moment condition, which in turn is handled using **(A2)** and **(A3)**. Finally, the uniform integrability follows from the Lemma as before, with **(A4)** replacing (17) and providing the required handle on  $|\phi_\psi(t)|$ .

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