

Three XOR-Lemmas — An Exposition

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Abstract. We provide an exposition of three lemmas that relate general properties of distributions over bit strings to the exclusive-or (xor) of values of certain bit locations.

The first XOR-Lemma, commonly attributed to Umesh Vazirani (1986), relates the statistical distance of a distribution from the uniform distribution over bit strings to the maximum bias of the xor of certain bit positions. The second XOR-Lemma, due to Umesh and Vijay Vazirani (*19th STOC*, 1987), is a computational analogue of the first. It relates the pseudorandomness of a distribution to the difficulty of predicting the xor of bits in particular or random positions. The third Lemma, due to Goldreich and Levin (*21st STOC*, 1989), relates the difficulty of retrieving a string and the unpredictability of the xor of random bit positions. The most notable XOR Lemma – that is the so-called Yao XOR Lemma – is not discussed here.

We focus on the proofs of the aforementioned three lemma. Our exposition deviates from the original proofs, yielding proofs that are believed to be simpler, of wider applicability, and establishing somewhat stronger quantitative results. Credits for these improved proofs are due to several researchers.

Keywords: Vector spaces, Kroniker and Fourier bases, pseudorandomness, space-bounded computation, one-way functions, hard-core predicates and functions.

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Preface

Unfortunately, the TCS community does not excell in its choice of names of various notions and phenomena. Consequently, we often find the same name used for several different issues. The name “XOR Lemma” is indeed a good example; at least four different technical statements are often referred to as XOR Lemmas. Indeed, the XOR operation features in each of these lemmas, but the actual context and contents of these lemmas vary.

The current article surveys three XOR lemmas, focusing on their proofs. As stated in the abstract, Yao’s XOR-Lemma is not one of the XOR Lemmas surveyed here; the interested reader is referred to [11].

1 The Information Theoretic XOR-Lemma

The Information Theoretic XOR-Lemma, commonly attributed to Umesh Vazirani, relates two measures of the “randomness” of distributions over n -bit long strings.

- The statistical difference from uniform; namely, the statistical difference (variation difference) between the “target” distribution and the uniform distribution over the set of all n -bit strings.
- The maximum bias of the xor of certain bit positions; namely, the bias of a 0-1 random variable obtained by taking the exclusive-or of certain bits in the “target” distribution.

It is well known that the statistical difference from uniform is bounded above by 2^n times the maximum bias of the xor’s. Several researchers have noticed that the factor in the bound can be improved to $\sqrt{2^n}$. We provide a four line proof of this fact. We also explain the reason for the popularity of the worse bound.

As motivation to the XOR-Lemma, we point out that it has been used in numerous works (e.g., Vazirani [20], Naor and Naor [16]). In a typical application, one first derives an upper bound on the maxbias of the constructed distribution, and then the XOR-Lemma is applied to infer an upper bound on the statistical difference from the uniform distribution.

Credit: The proof presented here has appeared as an appendix in [2].

1.1 Formal Setting

Let π be an arbitrary probability distribution over $\{0, 1\}^n$ and let μ denote the uniform distribution over $\{0, 1\}^n$ (i.e., $\mu(x) = 2^{-n}$ for every $x \in \{0, 1\}^n$). Let $x = x_1 \cdots x_n$ and $N \stackrel{\text{def}}{=} 2^n$. The XOR-Lemma relates two “measures of closeness” of π and μ .

- The statistical difference (“variation difference”) between π and μ ; namely,

$$\text{stat}(\pi) \stackrel{\text{def}}{=} \frac{1}{2} \cdot \sum_x |\pi(x) - \mu(x)| \quad (1)$$

- The “maximum bias” of the exclusive-or of certain bit positions in strings chosen according to the distribution π ; namely,

$$\text{maxbias}(\pi) \stackrel{\text{def}}{=} \max_{S \neq \emptyset} \{ |\pi(\{x : \bigoplus_{i \in S} x_i = 0\}) - \pi(\{x : \bigoplus_{i \in S} x_i = 1\})| \} \quad (2)$$

The XOR-Lemma, commonly attributed to Umesh Vazirani [20]¹, states that $\text{stat}(\pi) \leq N \cdot \text{maxbias}(\pi)$. The proof is based on viewing distributions as elements in an N -dimensional vector space and observing that the two measures

¹ The special case where the maxbias is zero appears in Chor et. al. [5]

considered by the lemma are merely two norms taken with respect to two different orthogonal bases (see Section 1.2). Hence, the XOR-Lemma follows from a (more general and quite straightforward) technical lemma that relates norms taken with respect to different orthonormal bases (see Section 1.3). It turns out this argument actually yields $\text{stat}(\pi) \leq \sqrt{N} \cdot \text{maxbias}(\pi)$, and it seems that the previously inferior bound of [20] was due to a less careful use of the same underlying ideas.

1.2 Preliminaries: the XOR-Lemma and vector spaces

Probability distributions over $\{0, 1\}^n$ are functions from $\{0, 1\}^n$ to the reals. Such functions form a N -dimensional vector space. We shall consider two alternative bases of this vector space.

The standard basis, denoted K , is the orthonormal basis defined by the **Kronker functions**; that is, the Boolean functions $\{k_\alpha : \alpha \in \{0, 1\}^n\}$, where $k_\alpha(x) = 1$ if $x = \alpha$. The statistical difference between two distributions equals (half) the norm L_1 of their difference taken in the above K basis.

On the other hand, the **maxbias** of a distribution equals the maximum **Fourier coefficient** of the distribution, which in turn corresponds to the max-norm (norm L_∞) of the distribution taken in a different basis. This basis is defined by the functions $\{b_S : S \subseteq \{1, 2, \dots, n\}\}$, where $b_S(x) = (-1)^{\sum_{i \in S} x_i}$. Note that $b_S(x) = 1$ if the exclusive-or of the bits $\{x_i : i \in S\}$ is 0 and $b_S(x) = -1$ otherwise. The new basis is orthogonal but not orthonormal. We hence consider the normalized basis, denoted F , consisting of the functions $f_S = \frac{1}{\sqrt{N}} \cdot b_S$.

Notation: Let B be an orthonormal basis and r an integer. We denote by $\mathbf{N}_r^B(v)$ the norm L_r of v with respect to the basis B . Namely, $\mathbf{N}_r^B(v) = (\sum_{e \in B} \langle e, v \rangle^r)^{1/r}$, where $\langle e, v \rangle$ is the absolute value of the inner product of the vectors e and v . We denote by $\mathbf{N}_\infty^B(v)$ the limit of $\mathbf{N}_r^B(v)$ when $r \rightarrow \infty$ (i.e., $\mathbf{N}_\infty^B(v)$ is $\max_{e \in B} \{\langle e, v \rangle\}$).

Clearly, $\text{stat}(\pi) = \frac{1}{2} \cdot \mathbf{N}_1^K(\pi - \mu)$ whereas $\text{maxbias}(\pi) = \sqrt{N} \cdot \mathbf{N}_\infty^F(\pi - \mu)$. Following is a proof of the second equality. Let $\delta(x) = \pi(x) - \mu(x)$. Clearly, $\text{maxbias}(\mu) = 0$ and hence $\text{maxbias}(\pi) = \text{maxbias}(\delta)$. Also $\sum_x \delta(x) = 0$. We get

$$\begin{aligned} \text{maxbias}(\delta) &= \max_{S \neq \emptyset} \{|\delta(\{x : b_S(x) = 1\}) - \delta(\{x : b_S(x) = -1\})|\} \\ &= \max_{S \neq \emptyset} \left\{ \left| \sum_x b_S(x) \cdot \delta(x) \right| \right\} \\ &= \sqrt{N} \cdot \max_S \left\{ \left| \sum_x f_S(x) \cdot \delta(x) \right| \right\} \\ &= \sqrt{N} \cdot \mathbf{N}_\infty^F(\delta) \end{aligned}$$

We now turn to the actual proof of the XOR Lemma.

1.3 Proof of the XOR-Lemma

The XOR-Lemma follows from the following technical lemma.

Technical Lemma: *For every two orthogonal bases A and B and every vector v , it holds that*

$$\mathbf{N}_1^A(v) \leq N \cdot \mathbf{N}_\infty^B(v). \quad (3)$$

This technical lemma has a three line proof:

For every orthogonal basis A ,

$$\mathbf{N}_1^A(v) \leq \sqrt{N} \cdot \mathbf{N}_2^A(v). \quad (4)$$

For every pair of orthonormal bases A and B ,

$$\mathbf{N}_2^A(v) = \mathbf{N}_2^B(v). \quad (5)$$

For every orthogonal basis B ,

$$\mathbf{N}_2^B(v) \leq \sqrt{N} \cdot \mathbf{N}_\infty^B(v) \quad (6)$$

Indeed, the Technical Lemma (i.e., Eq. (3)) is obtained by combining Eq. (4)–(6). Next, using this Technical Lemma, we get:

XOR-Lemma (revised): $\text{stat}(\pi) \leq \frac{1}{2} \cdot \sqrt{N} \cdot \text{maxbias}(\pi)$.

Proof: By the above

$$\text{stat}(\pi) = \frac{1}{2} \cdot \mathbf{N}_1^K(\pi - \mu) \leq \frac{1}{2} \cdot N \cdot \mathbf{N}_\infty^F(\pi - \mu) = \frac{1}{2} \cdot \sqrt{N} \cdot \text{maxbias}(\pi).$$

■

1.4 Discussion

The inferior bound, $\text{stat}(\pi) \leq N \cdot \text{maxbias}(\pi)$, has been derived by using one of the following two bounds instead of our Technical Lemma:

1. $\mathbf{N}_1^A(v) \leq \sqrt{N} \mathbf{N}_1^B(v) \leq \sqrt{N} \cdot N \mathbf{N}_\infty^B(v)$.

The first inequality is proved similarly to the proof of our Technical Lemma (i.e., using $\mathbf{N}_2^B(v) \leq \mathbf{N}_1^B(v)$ instead of Eq. (6)). The second inequality is trivial. Each of the two inequalities is tight, but their combination is wasteful.

2. $\mathbf{N}_1^A(v) \leq N \cdot \mathbf{N}_\infty^A(v) \leq N \cdot \sqrt{N} \mathbf{N}_\infty^B(v)$.

The second inequality is proved similarly to the proof of our Technical Lemma (i.e., using $\mathbf{N}_\infty^A(v) \leq \mathbf{N}_2^A(v)$ instead of Eq. (4)). The first inequality is trivial. Again, each of the inequalities is tight, but their combination is wasteful.

1.5 Variants

Using small variations on the foregoing argument, we obtain the following variants of the XOR-Lemma:

1. $\max_{x \in \{0,1\}^n} \{|\pi(x) - \mu(x)|\} \leq \text{maxbias}(\pi)$.
2. $\text{stat}(\pi) \leq \sqrt{\sum_{S \neq \emptyset} \text{bias}_S(\pi)^2}$, where $\text{bias}_S(\pi) = \sum_x b_S(x) \cdot \pi(x)$.

Proof: The first claim follows by using $\mathbf{N}_\infty^A(v) \leq \mathbf{N}_2^A(v)$ (instead of $\mathbf{N}_1^A(v) \leq \sqrt{N} \cdot \mathbf{N}_2^A(v)$), and obtaining $\mathbf{N}_\infty^K(\pi - \mu) \leq \sqrt{N} \cdot \mathbf{N}_\infty^F(\pi - \mu)$. The second claim follows by using $\mathbf{N}_1^A(v) \leq \sqrt{N} \cdot \mathbf{N}_2^B(v)$ and $\mathbf{N}_2^F(\pi - \mu) = \sqrt{\sum_{S \neq \emptyset} \text{bias}_S(\pi)^2}$. In both parts we also use $\text{bias}_\emptyset(\pi - \mu) = 0$. ■

1.6 Generalization to $\text{GF}(p)$, for any prime p

The entire treatment can be generalized to distributions over $\text{GF}(p)^n$, for any prime p . In this case, we redefine $N \stackrel{\text{def}}{=} p^n$, and let $\text{stat}(\pi)$ denote the statistical difference between π and the uniform distribution over $\text{GF}(p)^n$ (cf. Eq. (1)). Letting ω denote the p^{th} root of unity, we generalize Eq. (2) to

$$\text{maxbias}(\pi) \stackrel{\text{def}}{=} \max_{\beta \in \text{GF}(p)^n \setminus \{0\}^n} \left\{ \left| \sum_{e \in \text{GF}(p)} \omega^e \cdot \pi \left(\left\{ x : \sum_{i \in [n]} \beta_i x_i \equiv e \pmod{p} \right\} \right) \right| \right\}$$

The Fourier basis is generalized analogously: The new basis consists of the functions $\{b_\beta : \beta \in \text{GF}(p)^n\}$, where $b_\beta(x) = \omega^{\sum_{i \in [n]} \beta_i x_i}$. The normalized basis, denoted F , consists of the functions $f_\beta = N^{-1/2} \cdot b_\beta$.

Note that, in the case of $p = 2$, these definitions coincides with the definitions presented before. By following exactly the same manipulations as in the case of $p = 2$, we obtain the following generalization.

The XOR-Lemma, generalized to $\text{GF}(p)$: *Let π be an arbitrary distribution over $\text{GF}(p)^n$, and let μ denote the uniform distribution over $\text{GF}(p)^n$. Then*

1. $\text{stat}(\pi) \leq \frac{1}{2} \cdot \sqrt{N} \cdot \text{maxbias}(\pi)$.
2. $\max_{x \in \{0,1\}^n} \{|\pi(x) - \mu(x)|\} \leq \text{maxbias}(\pi)$.
3. $\text{stat}(\pi) \leq \frac{1}{2} \cdot \sqrt{\sum_{S \neq \emptyset} \text{bias}_S(\pi)^2}$, where $\text{bias}_S(\pi) = \sum_x b_S(x) \cdot \pi(x)$.

2 The Computational XOR-Lemma

We provide an exposition of the computational XOR-Lemma. By computational XOR-Lemma we refer to the assertion that a distribution on “short” strings is pseudorandom if and only if the xor of any of its bits is unpredictable. This Lemma was first proved by Umesh and Vijay Vazirani. The proof we present here is taken from the paper of Goldreich and Levin. We demonstrate the applicability of the computational XOR-Lemma by using it to construct pseudorandom generators with linear expansion factor that are “secure” against small (yet linear) bounded space machines.

2.1 Introduction

This section is concerned the relation between *two types of computationally restricted tests of randomness*. To be more precise, we are concerned with the pseudorandomness of a random variable Y given some partial information represented by an related random variable X . For sake of simplicity we write $X = f(R)$ and $Y = g(R)$ where f and g are fixed functions and R is a random variable uniformly distributed on strings of some length. Throughout this section, *we assume that f and g are polynomial-time computable*.

Tests of the first type are algorithms that, on input a pair (x, y) , output a single bit. We consider the probability that the test outputs 1 given that $x = f(r)$ and $y = g(r)$ where r is selected uniformly and compare it to the probability that the test outputs 1 given that $x = f(r)$ as before and y is selected (independently and) uniformly among the strings of length $|g(r)|$. We call the absolute value of the difference between these two probabilities the **distinguishing gap** of the test.

Tests of the second type are algorithms that, on input a string $f(r)$, output a single bit. The output is supposed to be the inner-product (mod 2) of the string $g(r)$ with some fixed string β (which is not all-zero). We consider the probability that the algorithm outputs the correct value given that r is selected uniformly. We call the absolute value of the difference between the success probability and the failure probability, the **advantage** of the algorithm. Note that the inner-product (mod 2) of $g(r)$ and β equals the exclusive-or of the bits in $g(r)$ that are located in positions corresponding to the 1 bits of β . Hence, tests of the second type try to predict the xor of bits in $g(r)$ that are in specified bit locations.

Vazirani and Vazirani [22] proved that if the tests are restricted to run in probabilistic polynomial-time and the length of $g(r)$ is logarithmic in the length of $f(r)$, then the two types of tests are equivalent in the following sense: There exists a test of the first type with a non-negligible distinguishing gap if and only if there exists a test of the second type with a non-negligible advantage². A different proof has appeared in Goldreich and Levin [10]. The interesting direction is, of course, the assertion that if there exists a test of the first type with a non-negligible distinguishing gap, then there exists a test of the second type with a non-negligible advantage³. This assertion is hereafter referred to as the **computational xor-lemma**.

The purpose of this section is to present a clear proof of the computational xor-lemma and to point out its applicability to other resource bounded machines. Our presentation follows the proof presented in [10], where all obvious details are omitted. Hence, the only advantage of our presentation is in its redundancy (w.r.t [10]).

² A function $\mu : \mathbf{N} \rightarrow \mathbf{R}$ is non-negligible if there exists a polynomial p such that for all sufficiently large n we have $\mu(n) > 1/p(n)$.

³ The opposite direction follows by noting that a test of a second type can be easily converted into a test of the first type: Just run the predicting algorithm and compare its outcome with the actual xor of the corresponding bits.

2.2 Proving the Computational XOR-Lemma

The proof proceeds via the counterpositive. That is, we show how to transform any test that distinguishes pairs $(f(r), g(r))$ from pairs $(f(r), y)$, where r and y are independently and uniformly distributed (among strings of adequate length), into a predictor of the xor of some bits of $g(r)$ from $f(r)$ such that the complexity of the predictor and its advantage are related to the complexity and distinguishing gap of the original tester. Actually, the construction yields a predictor that has a related advantage w.r.t a *random* subset of bits positions (rather than w.r.t *some* subset). The construction of the predictor and its analysis are captured by the following Technical Lemma.

In the following technical lemma, we present a particular algorithm, denoted G , that (given $f(r)$) tries to predict a specified xor of the bits of $g(r)$. The predictor G uses as subroutine a test, T , that (on input $f(r)$ and y) distinguishes a random y from $y = g(r)$. In particular, on input x and a subset S , the predictor selects y at random, runs the test T on inputs x and y , and output $\bigoplus_{i \in S} y_i$ if $T(x, y) = 1$ and the complement bit otherwise. The following lemma, lower bounds the advantage of the predictor G in terms of the distinguishing gap of the test T .

Technical Lemma (the core of the Computational XOR-Lemma): *Let f and g be arbitrary functions each mapping strings of the same length to strings of the same length. Let T be an algorithm (of the first type). Denote*

$$p \stackrel{\text{def}}{=} \Pr[T(f(r), g(r)) = 1] \quad (7)$$

and

$$q \stackrel{\text{def}}{=} \Pr[T(f(r), y) = 1], \quad (8)$$

where the probability is taken over all possible choices of $r \in \{0, 1\}^m$ and $y \in \{0, 1\}^{|g(r)|}$ with uniform probability distribution. Let G be an algorithm that, on input β and x , selects y uniformly in $\{0, 1\}^{|\beta|}$, and outputs $T(x, y) \oplus 1 \oplus (y, \beta)_2$, where $(y, \beta)_2$ is the inner product modulo 2 of y and β . Then,

$$\Pr[G(\beta, f(r)) = (g(r), \beta)_2] = \frac{1}{2} + \frac{p - q}{2^{|\beta|} - 1}, \quad (9)$$

where the probability is taken over all possible choices of $r \in \{0, 1\}^m$ and $\beta \in \{0, 1\}^{|g(r)|} \setminus \{0\}^{|g(r)|}$ with uniform probability distribution.

A full proof of the Technical Lemma is presented in Section 2.3. Before turning to that proof, we show that this lemma implies the Computational XOR-Lemma. This demonstration is immediate by the following two comments.

1. Algorithm G has almost the same complexities as T , with the exception that G must toss few more coins (to select β). Hence, G is randomized even in case T is deterministic.

2. Clearly, there exists a non-zero string β for which $\Pr[G(\beta, f(r)) = (g(r), \beta)_2] \geq \frac{1}{2} + \frac{p-q}{2^{|\beta|-1}}$, where the probability is taken over all possible choices of $r \in \{0, 1\}^m$ with uniform probability distribution. A string β with approximately such a performance can be found by sampling a string β and evaluating the performance of algorithm G with β as its first input. This requires ability to compute the functions f and g on many randomly selected instances (and collect the statistics). One should verify that this added complexity can be afforded. On the other hand, one should note that finding an appropriate β (i.e. on which G has almost the average advantage) may not be required (see the first remark below).

The following Computational XOR-Lemma follows as an immediate corollary to the Technical Lemma.

Computational XOR-Lemma: *Let \mathcal{C} be a class of randomized (or non-uniform) algorithms, such that \mathcal{C} is closed under sequential application of algorithms and contains an algorithm for computing $|g(r)|$ from $f(r)$. Suppose that every algorithm in the class \mathcal{C} , given $f(r)$, can predict the xor of a (given) random subset of the bits of $g(r)$ with (average) success probability at most $\frac{1}{2} + \epsilon$. Then, for every algorithm, T , in the class \mathcal{C} it holds that*

$$|\Pr[T(f(r), g(r)) = 1] - \Pr[T(f(r), y) = 1]| < 2^{|g(r)|} \cdot \epsilon$$

where r is selected uniformly in $\{0, 1\}^m$, the string y is selected uniformly and independently in $\{0, 1\}^{|g(r)|}$.

Remarks. As motivation to the Computational XOR-Lemma, we point out that it has been used in numerous works (e.g., Vazirani and Vazirani [22], Goldreich and Levin [10]). Another application of the Computational XOR-Lemma is presented in Section 2.4. In a typical application, the pseudorandomness of a short string is proved by showing that every xor of its bits is unpredictable (and using the Computational XOR-Lemma to argue that this suffices). Since it is typically the case that one can prove that the xor of a (given) *random* non-empty subset of the bits is unpredictable, the Computational XOR-Lemma can be used directly without finding an appropriate β (as suggested by a previous remark).

In case there are no computational restrictions on the tests, a stronger statement known as the XOR-Lemma can be proved: The statistical difference from uniform does not exceed $\sqrt{2^{|g(r)|}}$ times the maximum bias of a non-empty subset (see Section 1).

2.3 Proof of the Technical Lemma

Our goal here is to evaluate the success probability of algorithm G . In the following analysis we denote $\Pr_x[P(x, y)]$ the probability that $P(x, y)$ holds when x is distributed according to a distribution to be understood from the context, and y

is fixed. In the case that the predicate P depends on the test T , the probability will be taken also over the internal coin tosses of T . Hence, the coin tosses of T are implicit in the notation. In contrast, the additional coin tosses of G , namely the string y , are explicit in the notation. Hence, we rewrite

$$\begin{aligned} p &= \Pr_r[T(f(r), g(r)) = 1] \\ q &= \Pr_{r,y}[T(f(r), y) = 1] \end{aligned}$$

Recall that r is distributed uniformly on $\{0, 1\}^m$, whereas y is distributed uniformly on $\{0, 1\}^{|g(r)|}$. In the following analysis β is selected uniformly in $B \stackrel{\text{def}}{=} \{0, 1\}^{|g(r)| - 0^{|g(r)|}}$. Our aim is to evaluate $\Pr_{r,\beta,y}[G(\beta, f(r)) = (g(r), \beta)_2]$.

We start by fixing any $r \in \{0, 1\}^m$ and evaluating $\Pr_{\beta,y}[G(\beta, f(r)) = (g(r), \beta)_2]$. We define \equiv_β (resp., $\not\equiv_\beta$) such that $y \equiv_\beta z$ hold iff $(y, \beta)_2 = (z, \beta)_2$ (resp., $y \not\equiv_\beta z$ iff $(y, \beta)_2 \neq (z, \beta)_2$). We let $n \stackrel{\text{def}}{=} |g(r)|$. By the definition of G (i.e., $G(\beta, f(r)) = T(x, y) \oplus 1 \oplus (y, \beta)_2$, where $y \in \{0, 1\}^{|g(r)|}$ is uniformly selected by G) and elementary manipulations, we get

$$\begin{aligned} s_r &\stackrel{\text{def}}{=} \Pr_{\beta,y}[G(\beta, f(r)) = (g(r), \beta)_2] \\ &= \sum_{\beta \in B} \frac{1}{|B|} \cdot \Pr_y[G(\beta, f(r)) = (g(r), \beta)_2] \\ &= \frac{1}{|B|} \cdot \sum_{\beta \in B} \Pr_y[T(\beta, f(r)) = 1 \oplus (\beta, y)_2 \oplus (g(r), \beta)_2] \\ &= \frac{1}{2|B|} \cdot \sum_{\beta \in B} \Pr_y[T(f(r), y) = 1 | y \equiv_\beta g(r)] \\ &\quad + \frac{1}{2|B|} \cdot \sum_{\beta \in B} \Pr_y[T(f(r), y) = 0 | y \not\equiv_\beta g(r)] \\ &= \frac{1}{2} + \frac{1}{2|B|} \cdot \sum_{\beta \in B} \Pr_y[T(f(r), y) = 1 | y \equiv_\beta g(r)] \\ &\quad - \frac{1}{2|B|} \cdot \sum_{\beta \in B} \Pr_y[T(f(r), y) = 1 | y \not\equiv_\beta g(r)] \\ &= \frac{1}{2} + \frac{1}{2|B|} \cdot \frac{1}{2^{n-1}} \cdot \sum_{\beta \in B} \sum_{y \equiv_\beta g(r)} \Pr[T(f(r), y) = 1] \\ &\quad - \frac{1}{2|B|} \cdot \frac{1}{2^{n-1}} \cdot \sum_{\beta \in B} \sum_{y \not\equiv_\beta g(r)} \Pr[T(f(r), y) = 1] \\ &= \frac{1}{2} + \frac{1}{2^n \cdot |B|} \cdot \sum_y \sum_{\beta \in B \text{ s.t. } y \equiv_\beta g(r)} \Pr[T(f(r), y) = 1] \\ &\quad - \frac{1}{2^n \cdot |B|} \cdot \sum_y \sum_{\beta \in B \text{ s.t. } y \not\equiv_\beta g(r)} \Pr[T(f(r), y) = 1] \end{aligned}$$

Recall that $B = \{0, 1\}^n - 0^n$. Now, if $y \neq g(r)$ then the number of $\beta \in B$ for which $y \neq_{\beta} g(r)$ is 2^{n-1} (and the number of $\beta \in B$ for which $y \equiv_{\beta} g(r)$ is $2^{n-1} - 1$). On the other hand, if $y = g(r)$, then all $\beta \in B$ satisfy $y \equiv_{\beta} g(r)$. Hence, we get

$$\begin{aligned}
s_r - \frac{1}{2} &= \frac{1}{2^n |B|} \cdot \sum_{y \neq g(r)} ((2^{n-1} - 1) \cdot \Pr[T(f(r), y) = 1] - 2^{n-1} \cdot \Pr[T(f(r), y) = 1]) \\
&\quad + \frac{1}{2^n |B|} \cdot |B| \cdot \Pr[T(f(r), g(r)) = 1] \\
&= - \frac{1}{2^n |B|} \cdot \sum_{y \neq g(r)} \Pr[T(f(r), y) = 1] + \frac{1}{2^n |B|} \cdot |B| \cdot \Pr[T(f(r), g(r)) = 1] \\
&= - \frac{1}{|B|} \cdot \sum_y \frac{1}{2^n} \cdot \Pr[T(f(r), y) = 1] \\
&\quad + \frac{1}{2^n |B|} \cdot (|B| + 1) \cdot \Pr[T(f(r), g(r)) = 1] \\
&= - \frac{1}{|B|} \cdot \Pr_y[T(f(r), y) = 1] + \frac{1}{|B|} \cdot \Pr[T(f(r), g(r)) = 1]
\end{aligned}$$

Hence, for every r

$$\Pr_{\beta, y}[G(\beta, f(r)) = (g(r), \beta)_2] = \frac{1}{2} + \frac{\Pr[T(f(r), g(r)) = 1] - \Pr_y[T(f(r), y) = 1]}{|B|}$$

and so we have for uniformly chosen r

$$\Pr_{r, \beta, y}[G(\beta, f(r)) = (g(r), \beta)_2] = \frac{1}{2} + \frac{\Pr_r[T(f(r), g(r)) = 1] - \Pr_{r, y}[T(f(r), y) = 1]}{|B|}$$

and the lemma follows. ■

2.4 Application to pseudorandom generators for bounded space

We apply the Computational XOR-Lemma to the construction of pseudorandom generators with linear stretching that withstands tests of linearly bounded space. Namely, on input a random string of length n , the generator outputs a pseudorandom string of length cn withstanding tests of space en (where $e > 0$ is a constant depending on the constant $c > 1$). An alternative construction is immediate from the techniques presented by Nisan in [17].⁴ A third alternative construction was suggested by Noam Nisan (private communication) based on the ideas in [3].

The tests (or predictors) that we consider are non-uniform bounded-space machines with one-way access to the input (i.e., the string that they test). Hence, these machines can be represented by finite automata. By an $s(n)$ -space bounded

⁴ Use a constant number of hash functions.

machine we mean a finite automata with $2^{s(n)}$ states that is given an input of length n . For sake of simplicity, we sometimes discuss randomized automata. Clearly, randomness can be eliminated by introducing “more” non-uniformity.

Following is an overview of our construction. We begin by presenting a generator that extends seeds of length n into strings of length cn withstanding tests of space en , for a specific value of $c > 1$ (and $e > 0$). This generator is based on three observations:

1. Given two vectors, their inner-product mod 2 is unpredictable by machines of space significantly smaller than the length of these vectors.
2. With respect to such machines, the exclusive-or of bits resulting from the inner-product mod 2 of one vector and non-cyclic shifts of a second vector is also unpredictable. This holds because a machine predicting this exclusive-or can be transformed into a machine predicting the inner product of two vectors (cf. [10]).
3. Finally, using the computational XOR-Lemma, it follows that the bits resulting from the various inner-products are indistinguishable from random by space bounded machines.

The foregoing steps are detailed in Section 2.4.1. Next, in Section 2.4.2, we use this generator to construct, for every $k > 1$, a generator extending seeds of length n into strings of length $c^k \cdot n$ withstanding tests of space $(e/3)^k \cdot n$.

2.4.1 A construction for a specific expansion constant. The constants $c_1, \epsilon_1, c_0, \epsilon_0$ in the following construction and analysis will be determined in course of the analysis. In particular, $c_0 = \frac{1}{4}$, $\epsilon_0 = \frac{1}{6}$, $c_1 = 1 + \frac{c_0}{3}$, and $\epsilon_1 = \frac{\epsilon_0}{3}$, will do.

Construction 1: Using the notation $p_j(r_1 r_2 \cdots r_{2n}) \stackrel{\text{def}}{=} r_j r_{j+1} \cdots r_{j+n-1}$ and $b(x, s) \stackrel{\text{def}}{=} \sum_{i=1}^n x_i s_i \bmod 2$, consider the function $g: \{0, 1\}^{3n} \rightarrow \{0, 1\}^{c_0 n}$ defined by $g(x, r) = b(x, p_1(r)) \cdots b(x, p_{c_0 n}(r))$. Finally, consider the generator

$$g_1(x, r) = (x, r, g(x, r)). \quad (10)$$

This generator expands seeds of length $3n$ into strings of length $3n + c_0 n = c_1 \cdot 3n$. Clearly, the function g can be computed by an n -space machine. The robustness of the generator against $\epsilon_0 n$ -space machines follows from the following three claims.

Claim 1.1: Let A be an automaton with q states, and x, y be uniformly and independently selected in $\{0, 1\}^n$. Then

$$\Pr_{x,y}[A(x, y) = b(x, y)] \leq \frac{1}{2} + \sqrt{\frac{2q}{2^n}}.$$

Proof (adapted from [3]): By Lindsey Lemma (see [6, P. 88]), for every $X, Y \subseteq \{0, 1\}^n$, it holds that

$$\left| \sum_{x \in X, y \in Y} \frac{b(x, y)}{|X| \cdot |Y|} - \frac{1}{2} \right| \leq \sqrt{\frac{2^n}{|X| \cdot |Y|}} \quad (11)$$

Consider a partition of the set of all possible x 's according to the state in which the automaton is after reading x (i.e., the first half of its input), and denote the resulting sets X_1, X_2, \dots, X_q . Note that for every $x_1, x_2 \in X_j$ and every y , we have $A(x_1, y) = A(x_2, y)$. For each X_i , let Y_i^σ denote the sets of y 's for which $A(x, y) = \sigma$ given that $x \in X_i$. It follows that

$$\begin{aligned} \Delta &\stackrel{\text{def}}{=} \left| \Pr_{x, y}[A(x, y) = b(x, y)] - \frac{1}{2} \right| \\ &= \sum_{i=1}^q \sum_{\sigma \in \{0, 1\}} \Pr_{x, y}[x \in X_i \wedge y \in Y_i^\sigma] \cdot \left| \frac{|\{(x, y) \in X_i \times Y_i^\sigma : b(x, y) = \sigma\}|}{|X_i| \cdot |Y_i^\sigma|} - \frac{1}{2} \right| \\ &\leq \sum_{i=1}^q \sum_{\sigma \in \{0, 1\}} \Pr_{x, y}[x \in X_i \wedge y \in Y_i^\sigma] \cdot \sqrt{\frac{2^n}{|X_i| \cdot |Y_i^\sigma|}} \\ &= 2^{-3n/2} \cdot \sum_{i=1}^q \sum_{\sigma \in \{0, 1\}} \sqrt{|X_i| \cdot |Y_i^\sigma|} \\ &\leq 2^{-3n/2} \cdot \sqrt{2q} \cdot 2^n \end{aligned}$$

where the first inequality is due to Eq. (11) and the second inequality is due to (a special case of) the Cauchy-Schwartz Inequality.⁵ The claim follows. \square

Claim 1.2: Let $S \subseteq \{1, 2, \dots, m\}$, where $m < n$. Suppose that automaton A_S has q states and let

$$p \stackrel{\text{def}}{=} \Pr_{x, r} \left[A_S(x, r) = \bigoplus_{i \in S} b(x, p_i(r)) \right]$$

where the probability is taken over all random choices of $x \in \{0, 1\}^n$ and $r \in \{0, 1\}^{2^n}$. Then, there exists an automaton A with $q \cdot 2^{2m}$ states satisfying

$$\Pr_{x, y}[A(x, y) = b(x, y)] \geq p$$

where the probability is taken over all random choices of $x, y \in \{0, 1\}^n$.

Proof (adapted from [10]): Following is a construction of a randomized automaton A (randomization can be eliminated via non-uniformity). On input x, y , the predictor A produces a random string $r \in \{0, 1\}^{2|y|}$ satisfying $y_i =$

⁵ Specifically, we use $\sum_{j=1}^m \sqrt{a_j} \leq \sqrt{m \cdot \sum_{j=1}^m a_j}$.

$\sum_{j \in S} r_{i+j-1} \bmod 2$, for every $i \leq n$. This is done by setting the bits of r in increasing order such that r_k is randomly selected if either $k < t \stackrel{\text{def}}{=} \max(S)$ or $k \geq t+n$, and r_k is set to $y_{k-t+1} - \sum_{j \in S \setminus \{t\}} r_{k-t+j} \bmod 2$ for $k = t, t+1, \dots, t+n-1$. Hence, $\bigoplus_{j \in S} p_j(r) = y$, where $\bigoplus_{j \in S} v_j$ denotes the bit-by-bit exclusive or of the vectors v_j (where $j \in S$). The predictor A runs $A_S(x, r)$ and obtains a prediction for $\bigoplus_{j \in S} b(x, p_j(r)) = b(x, \bigoplus_{j \in S} p_j(r)) = b(x, y)$. The predictor uses at most $2m$ more space than G_S (for storing r), and the claim follows. \square

Claim 1.3: *For every automaton, T , with q states*

$$|\Pr[T(x, r, g(x, r)) = 1] - \Pr[T(x, r, y) = 1]| < 2^{|g(r)|} \cdot \sqrt{\frac{2q \cdot 2^{2c_0 n}}{2^n}}$$

where (x, r) is selected uniformly in $\{0, 1\}^{n+2n}$, the string y is selected uniformly in $\{0, 1\}^{|g(x, r)|}$.

Proof: Immediate by combining Claims 1.1 and 1.2 with the Computational XOR-Lemma. \square

Setting $c_0 = \frac{1}{4}$ and $\epsilon_1 = \frac{1}{6}$, we conclude that any $\epsilon_1 n$ -space bounded machine can distinguish $g_1(x, r)$ (where $xr \in \{0, 1\}^{3n}$) from a uniformly chosen string of length $(3 + c_0)n$ with gap at most $2^{-\epsilon_1 n}$. Hence, for constants $c_1 = 1 + \frac{1}{12}$ and $e_1 = \frac{1}{18}$, we have a generator extending strings of length n to strings of length $c_1 n$ so that no $\epsilon_1 n$ -space bounded machine can distinguish $g_1(s)$ (where $s = (x, r) \in \{0, 1\}^n$) from a uniformly chosen string of length $c_1 n$ with gap greater than $2^{-\epsilon_1 n}$. We say that g_1 has expansion factor c_1 and security constant e_1 .

2.4.2 Construction for any expansion constant. To achieve larger expansion we apply the generator again on small blocks of its output. This idea is taken from [9], but its usage in our context is restricted since in lower level the generator will be applied to shorter strings (and not to strings of the same length as done in [9]). The fact that in lower levels the generator is applied to shorter strings plays a key role in the proof that the resulting generator is indeed pseudorandom with respect to appropriate space-bounded machines.

In the sequel we show how to convert generators with expansion factor c into generators with expansion factor c^2 . Larger expansion factors are obtained by repeated application of the construction.

Construction 2: *Let g be a generator with expansion factor c and security constant e . We construct a generator g_2 with expansion factor c^2 and security constant $\frac{e^2}{3}$ as follows: $g_2(s) = g(r_1) \cdots g(r_t)$, where $r_1 \cdots r_t = g(s)$ such that $|r_j| = \frac{e}{2} \cdot |s|$ (for all $1 \leq j \leq t$) and $t = 2c/e$.*

To prove that the generator g_2 has security $\frac{e^2}{3}$ we consider a *hybrid* distribution H that results by selecting at random a string of length cn , partitioning it

into t blocks (each of length $\frac{c}{2}n$), and applying the generator g to each of them. First we show that H is hard to distinguish from random strings of length c^2n . Next, we show that H is hard to distinguish from the strings that g_2 generates on input a random seed of length n .

Claim 2.1 (indistinguishability of H and the uniform distribution): *Suppose that the automaton T has q states. Let $p_H \stackrel{\text{def}}{=} \Pr_{s_1 \dots s_t}[T(g(s_1) \dots g(s_t)) = 1]$ and $p_R \stackrel{\text{def}}{=} \Pr_{r_1 \dots r_t}[T(r_1 \dots r_t) = 1]$, where the probability is taken over all random choices of $s_1, \dots, s_t \in \{0, 1\}^{\frac{c}{2}n}$ and $r_1, \dots, r_t \in \{0, 1\}^{\frac{c}{2}n}$. Then, there exists an automaton T' with q states satisfying*

$$|\Pr_s[T'(g(s)) = 1] - \Pr_r[T'(r) = 1]| \geq \frac{|p_H - p_R|}{t}$$

where the probability is taken over all random choices of $s \in \{0, 1\}^{\frac{c}{2}n}$ and $r \in \{0, 1\}^{\frac{c}{2}n}$. Hence, if $q \leq \frac{c}{2}n$, then $|p_R - p_H| < \frac{1}{2} \cdot 2^{-\frac{c}{3}n}$.

Proof: For every $0 \leq i \leq t$, define

$$p_i \stackrel{\text{def}}{=} \Pr_{r_1 \dots r_i s_{i+1} \dots s_t}[T(r_1 \dots r_i g(s_{i+1}) \dots g(s_t)) = 1],$$

where the probability is taken over all random choices of $r_1, \dots, r_i \in \{0, 1\}^{\frac{c}{2}n}$ and $s_{i+1}, \dots, s_t \in \{0, 1\}^{\frac{c}{2}n}$. Namely, p_i is the probability that T outputs 1 on input taken from a hybrid distribution consisting of i “random” blocks and $t - i$ “pseudorandom” blocks. Clearly, $p_0 = p_H$ whereas $p_t = p_R$, and there exists $0 \leq i \leq t - 1$ such that $|p_i - p_{i+1}| \geq \frac{|p_0 - p_t|}{t}$. The test T' is obtained from T as follows. Fix a sequence $r_1, \dots, r_i \in \{0, 1\}^{\frac{c}{2}n}$ and $s_{i+2}, \dots, s_t \in \{0, 1\}^{\frac{c}{2}n}$ maximizing the distinguishing gap between the i^{th} and $i + 1^{\text{st}}$ hybrids. The starting state of test T' is the state to which T arrives on input r_1, \dots, r_i . The accepting states (i.e. states with output 1) of test T' are the state from which T reaches its accepting state when reading the string s_{i+2}, \dots, s_t . Clearly, T' has at most q states and distinguishes $r \in \{0, 1\}^{\frac{c}{2}n}$ from $g(s)$ (for $s \in \{0, 1\}^{\frac{c}{2}n}$) with gap at least $\frac{|p_H - p_R|}{t}$. Using the security hypothesis for g , it follows that $|p_R - p_H| < t \cdot 2^{-e \cdot \frac{c}{2}n} < \frac{1}{2} 2^{-\frac{c}{3}n}$ (for all sufficiently large n). The claim follows. \square

Note that the test constructed in the proof of Claim 2.1 examines strings of length $c \cdot \frac{c}{2}n$.

Claim 2.2 (indistinguishability of H and the output of g_2): *Suppose that the automaton T has q states and let $p_G \stackrel{\text{def}}{=} \Pr_s[T(g_2(s)) = 1]$ and $p_H \stackrel{\text{def}}{=} \Pr_{r_1 \dots r_t}[T(g(r_1) \dots g(r_t)) = 1]$, where the probability is taken over all random choices of $s \in \{0, 1\}^n$ and $r_1, \dots, r_t \in \{0, 1\}^{\frac{c}{2}n}$. Then, there exists an automaton T' with $q \cdot 2^{\frac{c}{2}n}$ states satisfying $|\Pr_s(T'(g(s)) = 1) - \Pr_r(T'(r) = 1)| \geq p_G - p_H$, where the probability is taken over all random choices of $s \in \{0, 1\}^n$ and $r \in \{0, 1\}^{cn}$. Hence, if $q \leq \frac{c}{2}n$, then $|p_G - p_H| < \frac{1}{2} 2^{-\frac{c}{3}n}$.*

Proof: The test T' is obtained from T as follows. On input $\alpha \in \{0, 1\}^{cn}$ (either random or pseudorandom), the test T' breaks α into t blocks, $\alpha_1, \dots, \alpha_t$, each of length $\frac{c}{2}n$. Then T' computes $\beta = \beta_1 \cdots \beta_t$ so that $\beta_i = g(\alpha_i)$, and applies T to the string β . (T' accepts α iff T accepts β .) If α is taken from the uniform distribution, then the resulting β is distributed according to H . On the other hand, if α is taken as the output of g on random seed s , then $\beta = g_2(s)$. The test T' distinguishes the above cases with gap $\geq |p_H - p_G|$, and can be implemented using $q \cdot 2^{\frac{c}{2}n}$ states. Using the security hypothesis for g , it follows that $|p_G - p_H| < 2^{-en} < \frac{1}{2}2^{-\frac{c^2}{3}n}$. The claim follows. \square

Note that the test constructed in the proof of Claim 2.2 evaluates g on strings of length $\frac{c}{2}n$. Combining Claims 2.1 and 2.2, we conclude that the generator g_2 has security constant $\frac{c^2}{3}$.

3 A Hard-Core Predicate for all One-Way Functions

A theorem of Goldreich and Levin [10] relates the following two computational tasks (regarding the function f). The first task is inverting a function f ; that is, given y , find an x so that $f(x) = y$. The second task is predicting, with non-negligible advantage, the exclusive-or of a subset of the bits of x when only given $f(x)$. More precisely, it has been proved that *if f cannot be efficiently inverted, then given $f(x)$ and r it is infeasible to predict the inner-product mod 2 of x and r better than obvious.*

We present an alternative proof to the original proof as appeared in [10]. The new proof, due to Charlie Rackoff, has two main advantages over the original one: It is simpler to explain and it provides better security (i.e., a more efficient reduction of inverting f to predicting the inner-product). The new proof was inspired by the proof in [1]. (We mention that the original proof provides a better starting point for the generalization presented in [12].)

3.1 Introduction

One-way functions are fundamental to many aspects of theory of computation. Loosely speaking, one-way are those functions that are easy to evaluate but hard to invert. However, many applications such as pseudorandom generators (see [4, 23] and [7, Chap. 3]) and secure encryption (see [13] and [8, Chap. 5]) require that the function has a “hard-core” predicate b . This value $b(x)$ should be easy to evaluate on input x , but hard to guess (with a noticeable correlation) when given only the value of $f(x)$. Intuitively, the hard-core predicate “concentrates” the one-wayness of the function in a strong sense. A natural question of practical and theoretical importance is *whether every one-way function has a hard-core predicate*. Prior to [10] only partial answers have been given:

1. Blum and Micali [4] proved that if the discrete exponentiation function is one-way, then it has a hard-core predicate.⁶ Analogous results for the RSA

⁶ This result was generalized to all Abelian groups in [14].

and Rabin functions (i.e. raising to a power and squaring modulo an integer, respectively) were obtained by Alexi, Chor, Goldreich, and Schnorr [1].

2. Yao [23] claimed that any one-way function f can be used to construct another one-way function f^* that has a hard-core predicate. The function f^* partitions its input into many shorter inputs and applies f to each of them in parallel (i.e., $f^*(x_1 \dots x_{k^3}) = f(x_1) \dots f(x_{k^3})$, where $|x_i| = k$). This claim was proved in [15] (see also [11]).

The drawback of the first set of results is that they are based on a particular intractability assumption (e.g. the hardness of the discrete logarithm problem). The second result constructs a predicate with security not bounded by a constant power of the security of f .

Goldreich and Levin [10] resolved the foregoing question by providing essentially every one-way function with a hard-core predicate (see Theorem 3 below). More specifically, for any time limit s (e.g. $s(n) = n$, or $s(n) = 2^{\sqrt{n}}$), the following tasks are equivalent for probabilistic algorithms running in time $s(|x|)^{O(1)}$:

1. Given $f(x)$ find x for at least a fraction $s(|x|)^{-O(1)}$ of the x 's.
2. Given $f(x)$ and p , $|p| = |x|$, guess the Boolean inner-product of x and p with a correlation (i.e. the difference between the success and failure probabilities) of $s(|x|)^{-O(1)}$.

We mention that, for any polynomial time computable f and b , the smallest (within a polynomial) such s exists and is called the **security** of f and b , respectively. The security is a constructible function, can be computed by trying all small guessing algorithms, and is assumed to grow very fast (at least $n^{1/o(1)}$).

3.2 Definitions

Loosely speaking, a *polynomial-time* function f is called **one-way** if any efficient algorithm can invert it only with negligible success probability. A *polynomial-time* predicate b is called a **hard-core** of a function f if all efficient algorithm, given $f(x)$, can guess $b(x)$ only with success probability that is negligibly better than half.

To simplify our exposition, we associate efficiency with polynomial-time and negligible functions as such decreasing smaller than $1/\text{poly}(n)$. By U_n we denote a random variable uniformly distributed over $\{0, 1\}^n$. For simplicity we consider only *length preserving* functions (i.e., $|f(x)| = |x|$ for every x).

Definition 1 (one-way function): *A one-way function, f , is a polynomial-time computable function such that for every probabilistic polynomial-time algorithm A' , every polynomial $p(\cdot)$, and all sufficiently large n it holds that*

$$\Pr[f(A'(Y_n)) = Y_n] < \frac{1}{2} + \frac{1}{p(n)}$$

where $Y_n = f(U_n)$.

Definition 2 (hard-core predicate): A polynomial-time computable predicate $b : \{0, 1\}^* \rightarrow \{0, 1\}$ is called a **hard-core** of a function f if for every probabilistic polynomial-time algorithm A' , every polynomial $p(\cdot)$, and all sufficiently large n it holds that

$$\Pr [A'(f(U_n)) = b(U_n)] < \frac{1}{2} + \frac{1}{p(n)}$$

3.3 The main result and its proof

The following result asserts that every one-way function has a closely related variant that has a hard-core predicate. The closely related variant (i.e., f' below) is obtained by padding the original function (i.e., f), and the security of f' is closely related to the security of f . Furthermore, the same hard-core predicate is used for all these variants (and is thus “universal” for them).

Theorem 3 (inner product mod 2 is an almost universal hard-core): Let f be an arbitrary one-way function, and let g be defined by $f'(x, r) \stackrel{\text{def}}{=} (f(x), r)$, where $|x| = |r|$. Let $b(x, r)$ denote the inner-product mod 2 of the binary vectors x and r . Then, the predicate b is a hard-core of the function g .

In other words, the theorem states that if f is one-way, then it is infeasible to guess the exclusive-or of a random subset of the bits of x when given $f(x)$ and the subset itself. We point out that f' maintains properties of f such as being length-preserving and being one-to-one. Furthermore, an analogous statement holds for collections of one-way functions with/without trapdoor etc. [7, Sec. 2.4].

As stated in Section 3.1, the proof of the foregoing Theorem 3 establishes a tight relation between the security of the one-way functions and the security of the corresponding hard-core predicate. This fact is one major advantage of Theorem 3 over Yao’s aforementioned construction [23].

3.3.1 The proof’s basic strategy. The proof uses a “reducibility argument” (see [7, Sec. 2.3.3]). Specifically, we assume (for contradiction) the existence of an efficient algorithm predicting the inner-product with advantage that is not negligible, and derive an algorithm that inverts f with related (i.e. not negligible) success probability. This contradicts the hypothesis that f is a one-way function. Thus, we show that inverting the function f is reduced to predicting $b(x, r)$ from $(f(x), r)$.

Let G be a (probabilistic polynomial-time) algorithm that on input $f(x)$ and r tries to predict the inner-product (mod 2) of x and r . Denote by $\varepsilon_G(n)$ the (overall) advantage of algorithm G in predicting $b(x, r)$ from $f(x)$ and r , where x and r are uniformly chosen in $\{0, 1\}^n$. That is,

$$\varepsilon_G(n) \stackrel{\text{def}}{=} \Pr [G(f(X_n), R_n) = b(X_n, R_n)] - \frac{1}{2}, \quad (12)$$

where here and in the sequel X_n and R_n denote two independent random variables, each uniformly distributed over $\{0, 1\}^n$. Assuming, towards the contradiction, that b is not a hard-core of f' means that exists an efficient algorithm G ,

a polynomial $p(\cdot)$ and an infinite set N so that for every $n \in N$ it holds that $\varepsilon_G(n) > \frac{1}{p(n)}$. We restrict our attention to this algorithm G and to n 's in this set N . In the sequel we shorthand ε_G by ε .

Our first observation is that, on at least an $\frac{\varepsilon(n)}{2}$ fraction of the x 's of length n , algorithm G has an $\frac{\varepsilon(n)}{2}$ advantage in predicting $b(x, R_n)$ from $f(x)$ and R_n . Namely,

Claim 3.1: *There exists a set $S_n \subseteq \{0, 1\}^n$ of cardinality at least $\frac{\varepsilon(n)}{2} \cdot 2^n$ such that for every $x \in S_n$, it holds that*

$$s(x) \stackrel{\text{def}}{=} \Pr[G(f(x), R_n) = b(x, R_n)] \geq \frac{1}{2} + \frac{\varepsilon(n)}{2}, \quad (13)$$

where here the probability is taken over all possible values of R_n and all internal coin tosses of algorithm G , whereas x is fixed.

Proof: The observation follows by an averaging argument. Namely, write $\text{Exp}[s(X_n)] = \frac{1}{2} + \varepsilon(n)$, and apply Markov Inequality. \square

In the sequel we restrict our attention to x 's in S_n . We will show an efficient algorithm that on every input y , with $y = f(x)$ and $x \in S_n$, finds x with very high probability. Contradiction to the one-wayness of f will follow by noting that $\Pr[U_n \in S_n] \geq \frac{\varepsilon(n)}{2}$.

Recall that $b(x, r) = \sum_{i=1}^n x_i r_i \pmod{2}$, where x_i (resp., r_i) denotes the i^{th} bit of x (resp., r). We highlight the fact that $b(x, r) \oplus b(x, s) = b(x, r \oplus s)$, which follows by $\sum_{i=1}^n x_i r_i + \sum_{i=1}^n x_i s_i \equiv \sum_{i=1}^n x_i (r_i \oplus s_i) \pmod{2}$.

3.3.2 A motivating discussion. Consider a fixed $x \in S_n$. By definition $s(x) \geq \frac{1}{2} + \frac{\varepsilon(n)}{2} > \frac{1}{2} + \frac{1}{2p(n)}$. Suppose, for a moment, that $s(x) > \frac{3}{4} + \frac{1}{2p(n)}$. In this case (i.e., of $s(x) > \frac{3}{4} + \frac{1}{\text{poly}(|x|)}$) retrieving x from $f(x)$ is quite easy. To retrieve the i^{th} bit of x , denoted x_i , we randomly select $r \in \{0, 1\}^n$, and compute $G(f(x), r)$ and $G(f(x), r \oplus e^i)$, where e^i is an n -dimensional binary vector with 1 in the i^{th} component and 0 in all the others, and $v \oplus u$ denotes the addition mod 2 of the binary vectors v and u . Clearly, if both $G(f(x), r) = b(x, r)$ and $G(f(x), r \oplus e^i) = b(x, r \oplus e^i)$, then

$$\begin{aligned} G(f(x), r) \oplus G(f(x), r \oplus e^i) &= b(x, r) \oplus b(x, r \oplus e^i) \\ &= b(x, e^i) \\ &= x_i \end{aligned}$$

(since $b(x, r) \oplus b(x, s) = b(x, r \oplus s)$). The probability that both equalities hold (i.e., both $G(f(x), r) = b(x, r)$ and $G(f(x), r \oplus e^i) = b(x, r \oplus e^i)$) is at least $1 - 2 \cdot (\frac{1}{4} - \frac{1}{\text{poly}(|x|)}) = \frac{1}{2} - \frac{1}{\text{poly}(|x|)}$. Hence, repeating the above procedure sufficiently many times and ruling by majority, we retrieve x_i with very high probability. Similarly, we can retrieve all the bits of x , and hence invert f on $f(x)$. However, the

entire analysis was conducted under (the unjustifiable) assumption that $s(x) > \frac{3}{4} + \frac{1}{2p(|x|)}$, whereas we only know that $s(x) > \frac{1}{2} + \frac{1}{2p(|x|)}$.

The problem with the above procedure is that it doubles the original error probability of algorithm G on inputs of form (x, \cdot) . Under the unrealistic assumption, that the G 's error on such inputs is significantly smaller than $\frac{1}{4}$, the “error-doubling” phenomenon raises no problems. However, in general (and even in the special case where G 's error is exactly $\frac{1}{4}$) the above procedure is unlikely to invert f . Note that the error probability of G can not be decreased by repeating G several times (e.g., G may always answer correctly on three quarters of the inputs, and always err on the remaining quarter). What is required is an *alternative way of using* the algorithm G , a way that does not double the original error probability of G .

The key idea is to generate the r 's in a way that requires applying algorithm G only once per each r (and x_i), instead of twice. The good news is that the error probability is no longer doubled, since we only need to use G to get an “estimate” of $b(x, r \oplus e^i)$. The bad news is that we still need to know $b(x, r)$, and it is not clear how we can know $b(x, r)$ without applying G . The answer is that we can guess $b(x, r)$ by ourselves. This is fine if we only need to guess $b(x, r)$ for one r (or logarithmically in $|x|$ many r 's), but the problem is that we need to know (and hence guess) $b(x, r)$ for polynomially many r 's. An obvious way of guessing these $b(x, r)$'s yields an exponentially vanishing success probability. The solution is to generate these polynomially many r 's so that, on one hand they are “sufficiently random” whereas on the other hand we can guess all the $b(x, r)$'s with non-negligible success probability. Specifically, generating the r 's in a *particular pairwise independent* manner will satisfy both (seemingly contradictory) requirements. We stress that in case we are successful (in our guesses for the $b(x, r)$'s), we can retrieve x with high probability. Hence, we retrieve x with non-negligible probability.

A word about the way in which the pairwise independent r 's are generated (and the corresponding $b(x, r)$'s are guessed) is indeed in place. To generate $m = \text{poly}(n)$ many r 's, we uniformly (and independently) select $l \stackrel{\text{def}}{=} \log_2(m + 1)$ strings in $\{0, 1\}^n$. Let us denote these strings by s^1, \dots, s^l . We then guess $b(x, s^1)$ through $b(x, s^l)$. Let us denote these guesses, which are uniformly (and independently) chosen in $\{0, 1\}$, by σ^1 through σ^l . Hence, the probability that all our guesses for the $b(x, s^i)$'s are correct is $2^{-l} = \frac{1}{\text{poly}(n)}$. The different r 's correspond to the different non-empty subsets of $\{1, 2, \dots, l\}$; that is, for every (non-empty) $J \subseteq \{1, 2, \dots, l\}$, we set $r^J \stackrel{\text{def}}{=} \bigoplus_{j \in J} s^j$. The reader can easily verify that the r^J 's are pairwise independent and each is uniformly distributed in $\{0, 1\}^n$ (see details below). The key observation is that

$$b(x, r^J) = b(x, \bigoplus_{j \in J} s^j) = \bigoplus_{j \in J} b(x, s^j). \quad (14)$$

Hence, our guess for the $b(x, r^J)$'s is $\bigoplus_{j \in J} \sigma^j$, and with non-negligible probability all our guesses are correct.

3.3.3 Back to the formal argument. Following is a formal description of the inverting algorithm, denoted A . We assume, for simplicity that f is length preserving (yet this assumption is not essential). On input y (supposedly in the range of f), algorithm A sets $n \stackrel{\text{def}}{=} |y|$, and $l \stackrel{\text{def}}{=} \lceil \log_2(2n \cdot p(n)^2 + 1) \rceil$, where $p(\cdot)$ is the polynomial guaranteed above (i.e., $\epsilon(n) > \frac{1}{p(n)}$ for the infinitely many n 's in N). Algorithm A uniformly and independently select $s^1, \dots, s^l \in \{0, 1\}^n$, and $\sigma^1, \dots, \sigma^l \in \{0, 1\}$. It then computes, for every non-empty set $J \subseteq \{1, 2, \dots, l\}$, a string $r^J \leftarrow \bigoplus_{j \in J} s^j$ and a bit $\rho^J \leftarrow \bigoplus_{j \in J} \sigma^j$. Next, for every $i \in \{1, \dots, n\}$ and every non-empty $J \subseteq \{1, \dots, l\}$, algorithm A computes $z_i^J \leftarrow \rho^J \oplus G(y, r^J \oplus e^i)$. Finally, algorithm A sets z_i to be the majority of the z_i^J values, and outputs $z = z_1 \cdots z_n$.

(Remark: In an alternative implementation of the foregoing ideas, the inverting algorithm, denoted A' , tries all possible values for $\sigma^1, \dots, \sigma^l$, and outputs only one of resulting strings z , with an obvious preference to a string z satisfying $f(z) = y$.)

Following is a detailed analysis of the success probability of algorithm A on inputs of the form $f(x)$, for $x \in S_n$, where $n \in N$. We start by showing that if the σ^j 's are correct, then, with constant probability, $z_i = x_i$ for all $i \in \{1, \dots, n\}$. This is proved by lower bounding the probability that the majority of the z_i^J 's equals x_i .

Claim 3.2: *For every $x \in S_n$ and every $i \in \{1, \dots, n\}$, it holds that*

$$\Pr \left[\left| \{J : b(x, r^J) \oplus G(f(x), r^J \oplus e^i) = x_i\} \right| > \frac{1}{2} \cdot (2^l - 1) \right] > 1 - \frac{1}{2n}$$

where $r^J \stackrel{\text{def}}{=} \bigoplus_{j \in J} s^j$ and the s^j 's are independently and uniformly chosen in $\{0, 1\}^n$.

Proof: For every J , define a 0-1 random variable ζ^J , so that ζ^J equals 1 if and only if $b(x, r^J) \oplus G(f(x), r^J \oplus e^i) = x_i$. The reader can easily verify that each r^J is uniformly distributed in $\{0, 1\}^n$. It follows that each ζ^J equals 1 with probability $s(x)$, which by $x \in S_n$, is at least $\frac{1}{2} + \frac{1}{2p(n)}$. We show that the ζ^J 's are pairwise independent by showing that the r^J 's are pairwise independent. For every $J \neq K$ we have, without loss of generality, $j \in J$ and $k \in K \setminus J$. Hence, for every $\alpha, \beta \in \{0, 1\}^n$, we have

$$\begin{aligned} \Pr [r^K = \beta \mid r^J = \alpha] &= \Pr [s^k = \beta \mid s^j = \alpha] \\ &= \Pr [s^k = \beta] \\ &= \Pr [r^K = \beta] \end{aligned}$$

and pairwise independence of the r^J 's follows. Let $m \stackrel{\text{def}}{=} 2^l - 1$. Using Chebyshev's Inequality, we get

$$\Pr \left[\sum_J \zeta^J \leq \frac{1}{2} \cdot m \right] \leq \Pr \left[\left| \sum_J \zeta^J - \left(\frac{1}{2} + \frac{1}{2p(n)} \right) \cdot m \right| \geq \frac{1}{2p(n)} \cdot m \right]$$

$$\begin{aligned}
&< \frac{\text{Var}(\zeta^{\{1\}})}{\left(\frac{1}{2p(n)}\right)^2 \cdot (2n \cdot p(n)^2)} \\
&< \frac{\frac{1}{4}}{\left(\frac{1}{2p(n)}\right)^2 \cdot 2n \cdot p(n)^2} \\
&= \frac{1}{2n}
\end{aligned}$$

The claim follows. \square

Recall that if $\sigma^j = b(x, s^j)$, for all j 's, then $\rho^J = b(x, r^J)$ for all non-empty J 's. In this case z output by algorithm A equals x , with probability at least one half. However, the first event happens with probability $2^{-l} = \frac{1}{2n \cdot p(n)^2}$ independently of the events analyzed in Claim 3.2. Hence, in case $x \in S_n$, algorithm A inverts f on $f(x)$ with probability at least $\frac{1}{4p(|x|)}$ (whereas the modified algorithm, A' , succeeds with probability at least $\frac{1}{2}$). Recalling that $|S_n| > \frac{1}{2p(n)} \cdot 2^n$, we conclude that, for every $n \in N$, algorithm A inverts f on $f(U_n)$ with probability at least $\frac{1}{8p(n)^2}$. Noting that A is polynomial-time (i.e., it merely invokes G for $2n \cdot p(n)^2 = \text{poly}(n)$ times in addition to making a polynomial amount of other computations), a contradiction to our hypothesis that f is one-way follows. The theorem follows. \blacksquare

3.3.4 Improving the efficiency of the inverting algorithm. In continuation to the proof of Theorem 3, we present guidelines for a more efficient inverting algorithm. In the sequel it will be more convenient to use the arithmetic of the reals instead of that of the Booleans. Hence, we denote $b'(x, r) = (-1)^{b(r, x)}$ and $G'(y, r) = (-1)^{G(y, r)}$.

1. Prove that, for every x , it holds that $\text{Exp}[b'(x, r) \cdot G'(f(x), r + e^i)] = s'(x) \cdot (-1)^{x_i}$, where $s'(x) \stackrel{\text{def}}{=} 2 \cdot (s(x) - \frac{1}{2})$.
2. Let v be an l -dimensional Boolean vector, and let R be a uniformly chosen l -by- n Boolean matrix. Prove that for every $v \neq u \in \{0, 1\}^l \setminus \{0\}^l$ it holds that vR and uR are pairwise independent and uniformly distributed in $\{0, 1\}^n$.
3. Prove that $b'(x, vR) = b'(xR^\top, v)$, for every $x \in \{0, 1\}^n$ and $v \in \{0, 1\}^l$.
4. Prove that, for every $x \in S_n$, with probability at least $\frac{1}{2}$ (over the choices of R as in Item 2), there exists $\sigma \in \{0, 1\}^l$ such that for every $1 \leq i \leq n$ the sign of $\sum_{v \in \{0, 1\}^l} b'(\sigma, v) G'(f(x), vR + e^i)$ equals the sign of $(-1)^{x_i}$. (Hint: $\sigma \stackrel{\text{def}}{=} xR^\top$.)
5. Let B be an 2^l -by- 2^l matrix with the (σ, v) -entry being $b'(\sigma, v)$, and let \bar{g}^i be an 2^l -dimensional vector with the v^{th} entry equal $G'(f(x), vR + e^i)$. Consider an inverting algorithm that computes $\bar{z}_i \leftarrow B \bar{g}^i$, for all i 's, and forms a matrix Z in which the columns are the \bar{z}_i 's. That is, the $(\sigma, i)^{\text{th}}$ entry in Z is $\sum_v b'(\sigma, v) \cdot G'(f(x), vR + e^i)$. The algorithm outputs a row of X such that applying f to it yields $f(x)$, where X is Boolean matrix such that its $(\sigma, i)^{\text{th}}$ entry is 1 iff the $(\sigma, i)^{\text{th}}$ entry in Z is negative.

- (a) Evaluate the success probability of this inverting algorithm.
- (b) Using the special structure of matrix B , show that the product Bg^i can be computed in time $l \cdot 2^l$.

Hint: B is the Sylvester matrix, which can be written recursively as

$$S_k = \begin{pmatrix} S_{k-1} & \overline{S_{k-1}} \\ S_{k-1} & S_{k-1} \end{pmatrix}$$

where $S_0 = +1$ and \overline{M} means flipping the $+1$ entries of M to -1 and vice versa.

3.4 Hard-Core Functions

We have just seen that every one-way function can be easily modified to have a hard-core predicate. In other words, the result establishes one bit of information about the preimage that is hard to approximate from the value of the function. A stronger result may say that several bits of information about the preimage are hard to approximate. For example, we may want to say that a specific pair of bits is hard to approximate, in the sense that it is infeasible to guess this pair with probability significantly larger than $\frac{1}{4}$. In general, a *polynomial-time* function, h , is called a hard-core of a function f if no efficient algorithm can distinguish $(f(x), h(x))$ from $(f(x), r)$, where r is a random string of length $|h(x)|$. We assume for simplicity that h is length regular (see below).

Definition 4 (hard-core function): *Let $h : \{0, 1\}^* \rightarrow \{0, 1\}^*$ be a polynomial-time computable function, satisfying $|h(x)| = |h(y)|$ for all $|x| = |y|$, and let $l(n) \stackrel{\text{def}}{=} |h(1^n)|$. The function $h : \{0, 1\}^* \rightarrow \{0, 1\}^*$ is called a **hard-core** of a function f if for every probabilistic polynomial-time algorithm D' , every polynomial $p(\cdot)$, and all sufficiently large n it holds that*

$$\left| \Pr [D'(f(X_n), h(X_n)) = 1] - \Pr [D'(f(X_n), R_{l(n)}) = 1] \right| < \frac{1}{p(n)}$$

where X_n and $R_{l(n)}$ are two independent random variables the first uniformly distributed over $\{0, 1\}^n$ and the second uniformly distributed over $\{0, 1\}^{l(n)}$.

Theorem 5 (almost universal hard-core functions): *Let f be an arbitrary one-way function, and let f_2 be defined by $f_2(x, s) \stackrel{\text{def}}{=} (f(x), s)$, where $|s| = 2|x|$. Let $c > 0$ be a constant, and $l(n) \stackrel{\text{def}}{=} \lceil c \log_2 n \rceil$. Let $b_i(x, s)$ denote the inner-product mod 2 of the binary vectors x and $(s_{i+1}, \dots, s_{i+n})$, where $s = (s_1, \dots, s_{2n})$. Then the function $h(x, s) \stackrel{\text{def}}{=} b_1(x, s) \cdots b_{l(|x|)}(x, s)$ is a hard-core of the function f_2 .*

The proof of the theorem follows by combining a proposition concerning the structure of the specific function h with a general lemma concerning hard-core functions. Loosely speaking, the proposition “reduces” the problem of approximating $b(x, r)$ given $f'(x, r)$ to the problem of approximating the exclusive-or

of any non-empty set of the bits of $h(x, s)$ given $f_2(x, s)$, where b and f' are the hard-core and the one-way function presented in Section 3.3. Since we know that the predicate $b(x, r)$ cannot be approximated from $f'(x, r)$, we conclude that no exclusive-or of the bits of $h(x, s)$ can be approximated from $f_2(x, s)$. The general lemma states that, for every “logarithmically shrinking” function h (i.e., h satisfying $|h(x)| = O(\log |x|)$), the function h is a hard-core of a function f if and only if the exclusive-or of any non-empty subset of the bits of h cannot be approximated from the value of f .

Proposition 6 (exclusive-ors of bits of h are hard-core predicates): *Let f , f_2 and b_i 's be as above. Let $I(n) \subseteq \{1, 2, \dots, l(n)\}$, $n \in \mathbf{N}$, be an arbitrary sequence of non-empty subsets, and let $b_{I(|x|)}(x, s) \stackrel{\text{def}}{=} \bigoplus_{i \in I(|x|)} b_i(x, s)$. Then, for every probabilistic polynomial-time algorithm A' , every polynomial $p(\cdot)$, and all sufficiently large n it holds that*

$$\Pr [A'(I(n), f_2(U_{3n})) = b_{I(n)}(U_{3n})] < \frac{1}{2} + \frac{1}{p(n)}$$

The proof is analogous to the proof of Claim 1.2 (presented in Section 2.4). Nevertheless, we detail the proof for sake of clarity.

Proof: The proof is by a “reducibility” argument. It is shown that the problem of approximating $b(X_n, R_n)$ given $(f(X_n), R_n)$ is reducible to the problem of approximating $b_{I(n)}(X_n, S_{2n})$ given $(f(X_n), S_{2n})$, where X_n , R_n and S_{2n} are independent random variable and the last is uniformly distributed over $\{0, 1\}^{2n}$. The underlying observation is that, for every $|s| = 2 \cdot |x|$,

$$b_I(x, s) = \bigoplus_{i \in I} b_i(x, s) = \bigoplus_{i \in I} b(x, \text{sub}_i(s)) = b(x, \bigoplus_{i \in I} \text{sub}_i(s))$$

where $\text{sub}_i(s_1, \dots, s_{2n}) \stackrel{\text{def}}{=} (s_{i+1}, \dots, s_{i+n})$. Furthermore, the reader can verify⁷ that for every non-empty $I \subseteq \{1, \dots, n\}$, the random variable $\bigoplus_{i \in I} \text{sub}_i(S_{2n})$ is uniformly distributed over $\{0, 1\}^n$, and that given a string $r \in \{0, 1\}^n$ and such a set I one can efficiently select a string uniformly in the set $\{s : \bigoplus_{i \in I} \text{sub}_i(s) = r\}$.

Now, assume to the contradiction, that there exists an efficient algorithm A' , a polynomial $p(\cdot)$, and an infinite sequence of sets (i.e., $I(n)$'s) and n 's such that

$$\Pr [A'(I(n), f_2(U_{3n})) = b_{I(n)}(U_{3n})] \geq \frac{1}{2} + \frac{1}{p(n)}$$

We first observe that for n 's satisfying the above inequality we can find in probabilistic polynomial time (in n) a set I satisfying

$$\Pr [A'(I, f_2(U_{3n})) = b_I(U_{3n})] \geq \frac{1}{2} + \frac{1}{2p(n)}$$

(i.e., by going over all possible I 's and experimenting with algorithm A' on each of them). Of course we may be wrong in these experiments, but the error probability can be made exponentially small.

⁷ Indeed, see the proof of Claim 1.2.

We now present an algorithm for approximating $b(x, r)$, from $y \stackrel{\text{def}}{=} f(x)$ and r . On input y and r , the algorithm first finds a set I as described above (this stage depends only on $|x|$, which equals $|r|$). Once I is found, the algorithm uniformly select a string s so that $\bigoplus_{i \in I} \text{sub}_i(s) = r$, and return $A'(y, s)$. Evaluation of the success probability of this algorithm is left as an exercise. ■

Lemma 7 (Computational XOR Lemma, revisited): *Let f and h be arbitrary length regular functions, and let $l(n) \stackrel{\text{def}}{=} |h(1^n)|$. Let D be an algorithm. Denote*

$$p \stackrel{\text{def}}{=} \Pr [D(f(X_n), h(X_n)) = 1] \quad \text{and} \quad q \stackrel{\text{def}}{=} \Pr [D(f(X_n), R_{l(n)}) = 1]$$

where X_n and R_l are as above. Let G be an algorithm that on input y and S (and $l(n)$), selects r uniformly in $\{0, 1\}^{l(n)}$, and outputs $D(y, r) \oplus 1 \oplus (\bigoplus_{i \in S} r_i)$, where $r = r_1 \cdots r_l$ and $r_i \in \{0, 1\}$. Then,

$$\Pr \left[G(f(X_n), I_l, l(n)) = \bigoplus_{i \in I_l} (h_i(X_n)) \right] = \frac{1}{2} + \frac{p - q}{2^{l(n)} - 1}$$

where I_l is a randomly chosen non-empty subset of $\{1, \dots, l(n)\}$ and $h_i(x)$ denotes the i^{th} bit of $h(x)$.

Proof: See Section 2. ■

It follows that, for logarithmically shrinking h 's, the existence of an efficient algorithm that distinguishes (with a gap that is not negligible in n) the random variables $(f(X_n), h(X_n))$ and $(f(X_n), R_{l(n)})$ implies the existence of an efficient algorithm that approximates the exclusive-or of a random non-empty subset of the bits of $h(X_n)$ from the value of $f(X_n)$ with an advantage that is not negligible.

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