1 Amplifying Success Probability

To amplify the success probability of Algorithm CountMin (in general case), we use median of independent repetitions (instead of minimum), and analyze it using the following (standard) concentration bounds.

**Chernoff-Hoeffding concentration bounds:** Let \( X = \sum_{i \in [n]} X_i \) where \( X_i \in [0, 1] \) for \( i \in [n] \) are independently distributed random variables. Then

\[
\forall t > 0, \quad \Pr[|X - \mathbb{E}[X]| \geq t] \leq 2e^{-2t^2/n}.
\]

\[
\forall 0 < \varepsilon \leq 1, \quad \Pr[X \leq (1 - \varepsilon) \mathbb{E}[X]] \leq e^{-\varepsilon^2 \mathbb{E}[X]/2}.
\]

\[
\forall 0 < \varepsilon \leq 1, \quad \Pr[X \geq (1 + \varepsilon) \mathbb{E}[X]] \leq e^{-\varepsilon^2 \mathbb{E}[X]/3}.
\]

\[
\forall t \geq 2e \mathbb{E}[X], \quad \Pr[X \geq t] \leq 2^{-t}.
\]

**Algorithm CountMin++:**

1. Run \( k = O(\log n) \) independent copies of algorithm CountMin, keeping in memory the vectors \( S^1, \ldots, S^k \) (and functions \( h^1, \ldots, h^k \))

2. Output: To estimate \( x_i \) report the median of all basic estimates \( \hat{x}_i = \text{median}\{\tilde{x}^l_i : l \in [k]\} \)

**Lemma:**

\[
\Pr[\hat{x}_i \in x_i \pm \alpha \|x\|_1] \leq 1/n^2.
\]

Proof: as seen in class, we define an indicator \( Y_l \) for the event that copy \( l \in [k] \) succeeds, then use one of the concentration bounds.

*These notes summarize the material covered in class, usually skipping proofs, details, examples and so forth, and possibly adding some remarks, or pointers. The exercises are for self-practice and need not be handed in. In the interest of brevity, most references and credits were omitted.*
**Theorem 1 [Cormode-Muthukrishnan, 2005]:** There is a streaming algorithm for $\ell_1$ point queries that uses a linear sketch of dimension $O(\alpha^{-1} \log n)$ (which implies that its memory requirement is this number of words) to achieve accuracy $\alpha \in (0, 1)$ with success probability $1 - 1/n^2$.

**Exer:** Use these concentration bounds to amplify the success probability of the algorithms we saw for Distinct Elements and for Probabilistic Counting (say from constant to $1 - 1/n^2$).

**Hint:** use independent repetitions + median.

### 2 $\ell_2$ Point Query via CountSketch

The idea is to hash coordinates to buckets (similar to algorithm CountMin), but furthermore use tug-of-war inside each bucket (as in algorithm AMS). The analysis will show it is a good estimate with error proportional to $\|x\|_2$ instead of $\|x\|_1$.

**Theorem 2 [Charikar, Chen and Farach-Colton, 2003]:** One can estimate $\ell_2$ point queries within error $\alpha$ with constant high probability, using a linear sketch of dimension $O(\alpha^{-2})$. It implies, in particular, a streaming algorithm.

It achieves better accuracy than CountMin ($\ell_2$ instead of $\ell_1$), but requires more storage ($1/\alpha^2$ instead of $1/\alpha$).

**Algorithm CountSketch:**

1. **Init:** Set $w = 4/\alpha^2$ and choose a hash function $h : [n] \rightarrow [w]$
2. Choose random signs $r_1, \ldots, r_n \in \{-1,+1\}$
3. **Update:** Maintain vector $S = [S_1, \ldots, S_w]$ where $S_j = \sum_{r(h(i)) = j} r_i x_i$.
4. **Output:** To estimate $x_i$ return $\tilde{x}_i = r_i \cdot S_{h(i)}$.

Storage requirement: $O(w) = O(\alpha^{-2})$ words, not counting storage of the random bits.

**Correctness:** We saw in class that $\Pr[|\tilde{x}_i - x_i|^2 \geq \alpha^2 \|x\|_2^2] \leq 1/4$, i.e., with high (constant) probability, $\tilde{x}_i \in x_i \pm \alpha \|x\|_2$.

**Exer:** Explain how to amplify the success probability to $1 - 1/n^2$ using the median of $O(\log n)$ independent copies.

### 3 Hash Functions

**Independent random variables:** Recall that two (discrete) random variables $X, Y$ are independent if

$$\forall x, y \quad \Pr[X = x, Y = y] = \Pr[X = x] \cdot \Pr[Y = y].$$

This is equivalent to saying that the conditioned random variable $X|Y$ has exactly the same distribution as $X$. It implies that in particular $\mathbb{E}[XY] = \mathbb{E}[X] \cdot \mathbb{E}[Y]$.  

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The above naturally extends to more than two variables, and then we say the random variables are mutually (or fully) independent.

**Pairwise independence:** A collection of random variables $X_1, \ldots, X_n$ is called *pairwise independent* if for all $i \neq j \in [n]$, the variables $X_i$ and $X_j$ are independent.

Example: Let $X, Y \in \{0, 1\}$ be random and independent bits, and let $Z = X \oplus Y$. Then $X, Y, Z$ are clearly not mutually (fully) independent, but they are pairwise independent.

Observation: When $X_1, \ldots, X_n$ are pairwise independent, the variance $\Var(\sum_i X_i)$ is exactly the same as if they were fully independent, because

$$
\Var(\sum_i X_i) = \mathbb{E}[\left(\sum_i X_i\right)^2] - \left(\mathbb{E}\left[\sum_i X_i\right]\right)^2 = \sum_{i,j} \mathbb{E}[X_i X_j] - \left(\sum_i \mathbb{E}[X_i]\right)^2.
$$

Consequently (and this is well-known): If $X_1, \ldots, X_n$ are pairwise independent (and have finite variance), then $\Var(\sum_i X_i) = \sum_i \Var(X_i)$.

The above definition extends to $k$-wise independence, where every subset of $k$ random variables should be independent.

**Pairwise independent hash family:** A family $H$ of hash functions $h : [n] \rightarrow [M]$ is called *pairwise independent* if for all $i \neq j \in [n]$,

$$
\forall x, y \in [M] \quad \Pr_{h \in H}\left[h(i) = x, h(j) = y\right] = \Pr[h(i) = x] \cdot \Pr[h(j) = y].
$$

This is the same as saying that $h(1), \ldots, h(n)$ are pairwise independent (when choosing random $h \in H$).

A common scenario is that each $h(i)$ is uniformly distributed over $[M]$.

**Universal hashing:** A family $H$ of hash functions $h : [n] \rightarrow [M]$ is called *2-universal* if for all $i \neq j \in [n]$,

$$
\Pr_{h \in H}\left[h(i) = h(j)\right] \leq 1/M.
$$

Observe that 2-universality is weaker than (follows from) pairwise independence when each $h(i)$ is distributed uniformly over $[M]$, but it suffices for many algorithms.